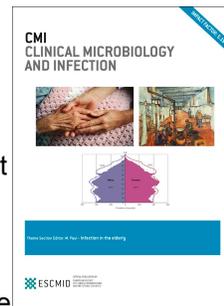


Journal Pre-proof

Outcomes of persons with COVID-19 in hospitals with and without standard treatment with (Hydroxy)chloroquine

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1 Original Article

2 Outcomes of Persons With COVID-19 in Hospitals With and
3 Without Standard Treatment With (Hydroxy)chloroquine
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100 Abstract

101

102 Objective

103 To compare survival of subjects with COVID-19 treated in hospitals that either did or did not
104 routinely treat patients with hydroxychloroquine or chloroquine.

105 Methods

106 We analysed data of COVID-19 patients treated in 9 hospitals in the Netherlands. Inclusion dates
107 ranged from February 27th 2020, to May 15th, when the Dutch national guidelines no longer
108 supported the use of (hydroxy)chloroquine. Seven hospitals routinely treated subjects with
109 (hydroxy)chloroquine, two hospitals did not. Primary outcome was 21-day all-cause mortality. We
110 performed a survival analysis using log-rank test and Cox-regression with adjustment for age, sex and
111 covariates based on premorbid health, disease severity, and the use of steroids for adult respiratory
112 distress syndrome, including dexamethasone.

113 Results

114 Among 1949 included subjects, 21-day mortality was 21.5% in 1596 subjects treated in hospitals that
115 routinely prescribed (hydroxy)chloroquine, and 15.0% in 353 subjects that were treated in hospitals
116 that did not. In the adjusted Cox-regression models this difference disappeared, with an adjusted
117 hazard ratio of 1.09 (95%CI 0.81-1.47). When stratified by actually received treatment in individual
118 subjects, the use of (hydroxy)chloroquine was associated with an increased 21-day mortality (HR
119 1.58; 95%CI 1.24-2.02) in the full model.

120 Conclusions

121 After adjustment for confounders, mortality was not significantly different in hospitals that routinely
122 treated patients with (hydroxy)chloroquine, compared with hospitals that did not. We compared
123 outcomes of hospital strategies rather than outcomes of individual patients to reduce the chance of
124 indication bias. This study adds evidence against the use of (hydroxy)chloroquine in hospitalised
125 patients with COVID-19.

126 Introduction

127

128 The spread of SARS-CoV-2, leading to the current pandemic of COVID-19, has a profound global
129 impact on daily life, morbidity and mortality. Several preliminary studies have reported that the
130 antimalarial agents hydroxychloroquine and chloroquine, or (H)CQ, alone or in combination with the
131 antibiotic azithromycin, can have a suppressive effect on the viral replication, and might decrease the
132 mortality of COVID-19¹⁻⁵. So far, clinical studies have been hampered by confounding by
133 indication^{1,2,4,5}, monocentre setup^{2,3}, and small numbers of included subjects³. A recently published
134 systematic review⁶, a published randomized controlled trial⁷ and an RCT only available in pre-print⁸,
135 suggested that hydroxychloroquine is not effective in patients admitted to hospital. Side effects of
136 (H)CQ are well-known, and include fever and cardiac arrhythmias. While we are awaiting definite
137 results from more RCTs, cohort studies can provide quick closure of existing knowledge gaps. When
138 treatment assignment in cohort studies is based on prescriber discretion, the risk of indication bias
139 (even after covariate adjustment) remains high. However, our database of Dutch hospitals contains
140 data of subjects from hospitals that either routinely prescribed (H)CQ or did not prescribe it at all,
141 offering a unique opportunity to compare both strategies. The comparison of different treatment
142 strategies among hospitals leads to a significant reduction of (indication) bias. The objective of this
143 study was to compare the effect of hospital-wide COVID-19 treatment strategies with or without
144 routine (H)CQ use on all-cause 21-day mortality.

145

146

147

148 **Methods**

149

150 We used data from the ongoing CovidPredict Clinical Course Cohort containing over 2,000 persons
151 with COVID-19⁹, from 9 hospitals in the Netherlands, including two university hospitals. Included in
152 the database were all subjects admitted to hospital with positive SARS-CoV-2 PCR of nasopharynx,
153 throat, sputum or bronchoalveolar lavage samples, or CT-scan abnormalities that were typical for
154 COVID-19 (CO-RADS 4 and 5)¹⁰, without another explanation for the abnormalities than COVID-19.
155 Inclusion dates ranged from the first admitted case in the Netherlands on February 27th 2020, to May
156 15th, when the Dutch national guidelines no longer advised the use of (H)CQ. We excluded patients <
157 18 years and patients who were transferred to or from another hospital. Dosage of chloroquine base
158 was: loading dose of 600 mg, followed by 300 mg twice a day for a total of 5 days. Dosage of
159 hydroxychloroquine sulphate was 400 mg twice daily on the first day, followed by 200 mg twice daily
160 on days 2 to 5. Among the seven (H)CQ-hospitals, the timing of start of (H)CQ treatment differed;
161 three hospitals started at the moment of COVID-19 diagnosis, four started after diagnosis but only
162 when patients clinically deteriorated e.g., when there was an increase in respiratory rate or increase
163 in use of supplemental oxygen. The two hospitals that did not routinely treat subjects with (H)CQ
164 (i.e., the non-(H)CQ-hospitals), offered best supportive care, including oxygen therapy and
165 potentially antibiotic therapy, according to local guidelines and prescriber discretion. Participating
166 hospitals did not routinely prescribe other experimental medication (e.g., lopinavir/ritonavir,
167 remdesivir or steroids, see Table 1). Subjects who were incidentally treated with these drugs were
168 included in the study. Primary outcome was 21-day all-cause mortality, defined as hospital mortality,
169 or discharge to a hospice care facility. A waiver for the use of hospital record data was obtained
170 through the Institutional Review Board of Amsterdam UMC; however, patients were given the
171 opportunity to opt out. We collected data according to the collection protocol of the World Health
172 Organization. Missing covariates were imputed using multiple imputation with the MICE package
173 (version 3.8.0) and the outcomes were determined by pooling the results of 25 imputed datasets ¹¹.

174 We performed a regression analyses and determined the pooled effect. Missing data range for all
175 covariates was less than 2.8%, except for obesity (missing data 6.2%) and use of corticosteroids
176 (22.3%). In the primary analysis, we compared effectiveness of (H)CQ versus non-(H)CQ hospital
177 strategies, irrespective of actual individual (H)CQ treatment. We performed a survival analysis using
178 log-rank test and Cox-regression with adjustment for age, sex, time in the pandemic (i.e., the number
179 of elapsed days after March 1st 2020 at hospital admission), and covariates based on premorbid
180 health (i.e., history of lung, kidney and cardiovascular disease, diabetes mellitus, obesity, and
181 neoplasms or hematologic disease), disease severity during presentation (respiratory rate, oxygen
182 saturation) and the use of steroids, including dexamethasone, for adult respiratory distress
183 syndrome (ARDS)^{12,13}. We repeated the analyses comparing actually received treatment, with (H)CQ.
184 In a secondary analysis, we used a composite endpoint (either mechanical ventilation or all-cause
185 mortality) at 21 days. As a sensitivity analysis, we performed a complete case analysis using inverse
186 probability weighting of propensity scores (determined using the same covariates). We performed a
187 subgroup analysis in (H)CQ hospitals that started (H)CQ directly from the moment of diagnosis versus
188 outcomes in non-(H)CQ hospitals. All statistical analyses were performed using R versions 3.6.3 (R
189 Foundation, Vienna, Austria).

190

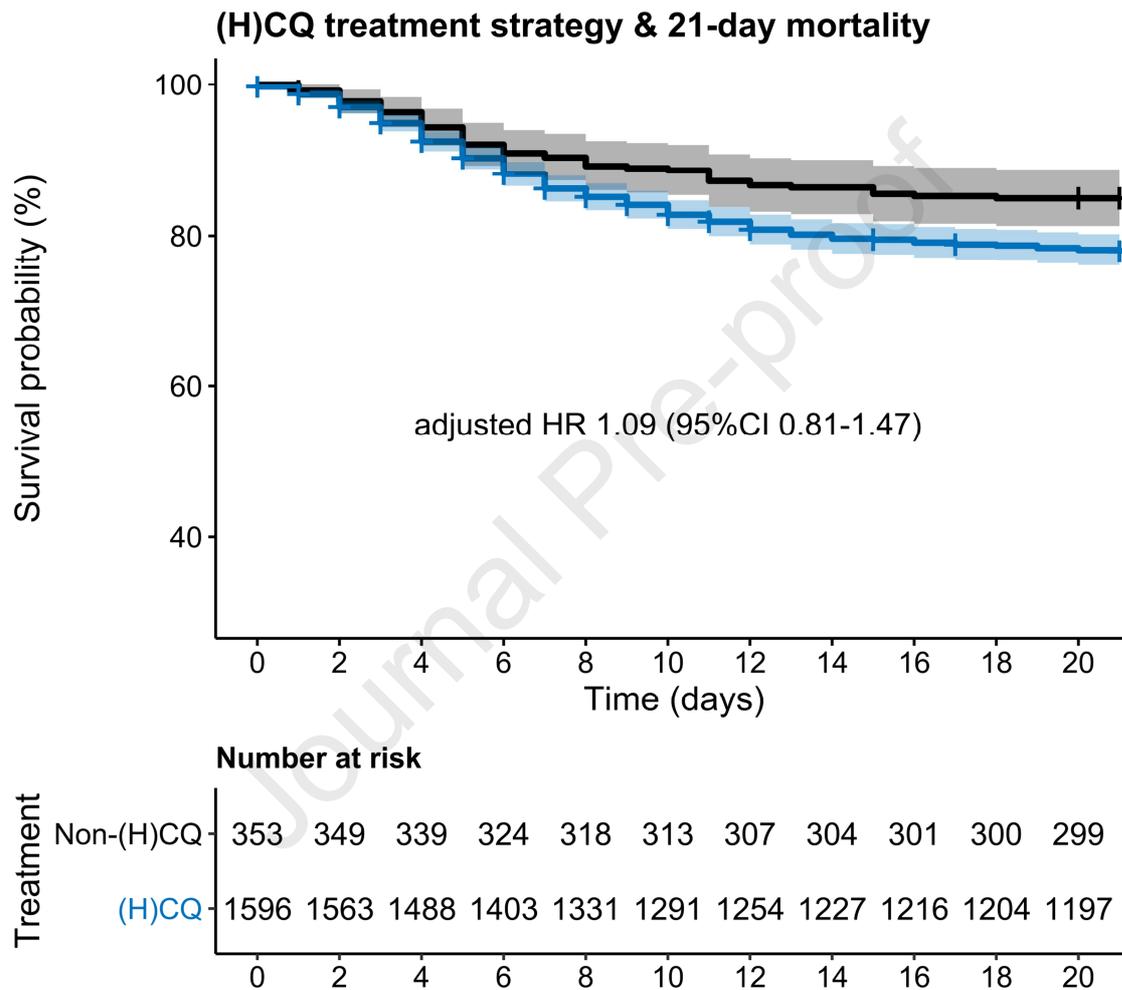
191 Results

192

193 We analysed results of 1949 of 2152 subjects admitted before May 15th 2020. 203 were excluded
194 because they were transferred from another hospital. No subject opted out. Demographic data are
195 shown in Table 1. Follow-up data were missing for 20 (1.0%) subjects. The patients with missing
196 outcome data were included Table 1 and in the survival analysis, and were censored at the last day at
197 which clinical information was available in the database. In total, 1596 subjects were treated in
198 hospitals where (H)CQ was a standard part of treatment strategy ((H)CQ hospitals) and 353 in

199 hospitals where (H)CQ was not a standard part of treatment (non-(H)CQ hospitals). The two non-
200 (H)CQ hospitals were both university hospitals. In (H)CQ-hospitals, 54.7% of the subjects received
201 (H)CQ, compared with 2.0% of the subjects in the non-(H)CQ-hospitals. In (H)CQ hospitals that
202 routinely starting (H)CQ at the moment of COVID-19 diagnosis, 48.3% of subjects received (H)CQ, in
203 hospitals that started (H)CQ at clinical deterioration, 61.9% received (H)CQ. Among the seven (H)CQ-
204 hospitals, two used hydroxychloroquine during the first half and chloroquine during the second half
205 of the epidemic, whereas five hospitals used chloroquine only. Subjects in (H)CQ-hospitals were
206 older (68 (SD: 14) vs 62 (SD: 15) years) and had a higher prevalence of chronic pulmonary disease
207 (27.7 vs 22.1) than subjects in the non-(H)CQ-hospitals. Respiratory rate and peripheral oxygen
208 saturation during admission were similar in both hospital groups (see Table 1). In (H)CQ-hospitals,
209 9.6% of subjects received corticosteroids for ARDS and 4.0% were in a study protocol of an
210 experimental SARS-CoV-2 directed antiviral (e.g., lopinavir/ritonavir) or immunomodulatory drug trial
211 (e.g., imatinib, anti-complement C5), versus 2.3% and 11.3% in non-(H)CQ-hospitals, respectively.
212 Figure 1 shows the survival of subjects in (H)CQ- versus non-(H)CQ-hospitals. Unadjusted mortality at
213 day 21 was significantly higher in the (H)CQ hospitals (343/1596, 21.5%) compared with the non
214 (H)CQ-hospitals (53/353, 15.0%, $p=0.008$). However, in the Cox-regression models, this difference
215 disappeared, with an adjusted hazard ratio of 1.09 (95%CI 0.81-1.47, Figure 1, Table 2). When
216 stratified by actually received treatment, the use of (H)CQ was associated with an increased 21-day
217 mortality (HR 1.58; 95%CI 1.24-2.02, Table 3) in the full model. In the secondary analysis with either
218 mechanical ventilation or all-cause mortality at 21 days, there were no statistically significant
219 differences between the (H)CQ and non-(H)CQ hospitals (crude $p=0.055$, adjusted HR 0.87 (95%CI
220 0.68-1.10), Online Supplement 1). The complete analysis using propensity scores for treatment
221 strategy and actual treatment showed similar results (see Table 4). An overview of the distribution of
222 the propensity scores is given in Online Supplement 2¹⁴. The sensitivity analysis of hospitals routinely
223 starting (H)CQ treatment from the moment of COVID-19 diagnosis (i.e., (H)CQ hospitals without the
224 hospitals that initiated (H)CQ treatment upon clinical deterioration) compared with non-(H)CQ-

225 hospitals, showed similar results with a significantly higher unadjusted 21-day mortality in (H)CQ-
 226 hospitals (154/670, 23.0%), compared with non-(H)CQ hospitals (53/353, 15.0%, $p=0.002$). This was
 227 attenuated towards a HR of 0.98 (95%CI 0.70-1.37) after adjustment for age, sex, comorbidities, and
 228 disease severity at presentation (Online Supplement 3).
 229



230
 231 Figure 1: Kaplan-Meier analysis of 21 day mortality of subjects in the (H)CQ-hospitals (blue) versus
 232 non-(H)CQ-hospitals (black), showing a significantly higher 21-day mortality in (H)CQ hospitals,
 233 $p=0.004$. This was attenuated towards a HR of 1.09 (95%CI 0.81-1.47) in the full regression model,
 234 see Table 2. Shaded areas indicate 95% confidence interval.

235

236

237 **Discussion**

238

239 Mortality in subjects treated in hospitals that routinely prescribed (H)CQ was not significantly
240 different from those treated in hospitals that routinely did not prescribe (H)CQ after adjustment for
241 age, sex, medical history, disease severity at presentation and steroid use during treatment.

242 Similarly, we found an increased risk of death among subjects who had actually received treatment
243 with (H)CQ, which has likely been driven by indication bias, as in four of the seven (H)CQ-hospitals,
244 (H)CQ was only prescribed upon clinical deterioration. The unique characteristics of our study cohort
245 enabled a study design that minimized indication bias. Our results add further weight to existing
246 evidence against the use of (H)CQ for the treatment COVID-19.

247

248 The strength of this study is that data were collected in nine hospitals, including two university
249 hospitals, in the Netherlands during the COVID-19 epidemic. Data collection was set up prospectively
250 and the database included data on all consecutive subjects admitted to general medicine and
251 pulmonology wards, and to intensive care units. The database was set up according to the WHO
252 standards, which enabled data comparison and uniformity of data among the different participating
253 centres. The comparison of hospital-defined treatment strategies rather than the treatment actually
254 received led to a lower risk of indication bias compared with previous studies^{1,2,4,5}. We roughly
255 estimate the extend of the effect of indication bias to be the difference in outcome between the
256 uncorrected and the corrected model. Further strengths include the multicentre setup^{2,3}, as
257 mentioned above, and the relatively large numbers of included subjects³.

258

259 There are some limitations we need to address. Although health care in the Netherlands has a
260 homogeneous setup, there was some variability in standard protocols among the hospitals that could

261 have led to residual confounding. The two non-(H)CQ-hospitals were tertiary (university) centres,
262 whereas the (H)CQ-hospitals comprised both secondary and tertiary care hospitals. Before the
263 COVID-19 pandemic, the tertiary care hospitals and their intensive care units function as referral
264 centres for local secondary care hospitals. Since we excluded subjects transferred to and from other
265 hospitals, the referral role of the tertiary care hospitals, including the university hospitals, was
266 minimized. Furthermore, subjects in the (H)CQ hospitals were more likely to receive steroid
267 treatment, while subjects in the non-(H)CQ hospitals were more likely to receive other experimental
268 immunomodulatory drugs. The numbers of the individual types of medication were small, making it
269 impossible to draw conclusions from these differences. The results of the RECOVERY trial, suggested
270 a lower mortality in patients treated with dexamethasone¹⁵. Treatment with dexamethasone could
271 therefore have resulted in a lower mortality in the group of (H)CQ hospitals. We did not find such an
272 effect, even after correction in the full model. We also used extensive covariate adjustments, using
273 various methods to minimize influence of differences in patient population among hospitals, and the
274 similarity in outcomes between these methods is reassuring in this regard. Finally, because not every
275 subject in the (H)CQ-hospitals actually received (H)CQ, the current efficacy estimate in our study is
276 likely an underestimation of the true (H)CQ effect. Performing an instrumental variable analysis
277 would have provided an approximation of this true effect, but because the current efficacy point
278 estimates point toward harm rather than benefit of (H)CQ, this likely would not have changed our
279 conclusions.¹⁶

280

281 Despite the positive results of some studies resulting in widespread use of (H)CQ, our study did not
282 show a benefit of (H)CQ treatment. This may be explained by the timing of the administration of the
283 drug and its specific working mechanism. Chloroquine binds *in silico* and *in vitro* with high affinity to
284 sialic acids and gangliosides of SARS-CoV-2. These bindings inhibit the interaction at non-toxic plasma
285 levels with ACE-2 receptors and could hypothetically stop the cascade from formation of pulmonary
286 infiltrations to full blown ARDS and death¹⁷⁻¹⁹. The antiviral activity might be more effective in the

287 pre-clinical setting as the deterioration in the hospital is more an effect of the cytokine storm
288 provoked by SARS-CoV-2 than an effect of the viral infection itself. This hypothesis might explain why
289 the clinical benefit for admitted subjects was absent in our study, although we did not observe a
290 difference in outcome among subjects treated early (at diagnosis) and among those treated later
291 upon clinical deterioration.

292

293 Our results are in line with recently published studies. A RCT suggest a similar lack of effect of
294 hydroxychloroquine with higher rate of adverse effect than in supportive care⁷. Another RCT,
295 published in preprint only, suggested a higher mortality in patients treated with hydroxychloroquine
296 compared with those treated with supportive care⁸. Given the current evidence, we would argue
297 against the use of (H)CQ in hospitals outside the setting of randomized controlled clinical trials.

298

299

300 **Conflict of Interest Disclosure**

301 The authors do not have any relevant financial or other disclosures.

302

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309 and interpretation of the data.

310

311

312 **Contribution of authors**

313 All authors have made substantial contributions to the following: (1) the conception and design of
314 the study, or acquisition of data, or analysis and interpretation of data, (2) drafting the article or
315 revising it critically for important intellectual content, (3) final approval of the version to be
316 submitted.

317 Table 1: Baseline characteristics

318

	Overall	Non-(H)CQ hospital	(H)CQ hospitals
N	1949	353	1596
Age (mean (SD))	66.71 (14.60)	62.02 (15.14)	67.75 (14.28)
Women (%)	771 (39.6)	155 (43.9)	616 (38.6)
Chronic cardiac disease (%)	587 (30.7)	75 (21.3)	512 (32.8)
Hypertension (%)	915 (47.6)	162 (46.2)	753 (47.9)
Asthma or chronic pulmonary disease (%)	510 (26.7)	78 (22.1)	432 (27.7)
Chronic kidney disease (%)	221 (11.6)	38 (10.8)	183 (11.8)
Diabetes (%)	501 (26.4)	96 (27.2)	405 (26.2)
Malignancy or chronic hematologic disorder (%)	194 (10.2)	44 (12.5)	150 (9.6)
Smoking (%)	92 (6.2)	18 (6.3)	74 (6.2)
Obesity (%)	556 (30.4)	107 (35.3)	449 (29.4)
Use of (H)CQ	648 (42.6)	7 (2.0)	641 (54.7)
Use of steroids for ARDS (%)	120 (7.9)	8 (2.3)	112 (9.6)
Participation in drug trial (%)	85 (5.7)	39 (11.3)	46 (4.0)
Respiratory rate (mean (SD))	23.20 (6.94)	24.29 (7.32)	22.95 (6.83)
Temperature, °C, (median [IQR])	37.80 [37.00, 38.60]	37.30 [36.50, 38.20]	38.00 [37.10, 38.70]
Peripheral oxygen saturation, %, (median [IQR])	94.00 [91.00, 96.00]	95.00 [91.00, 97.00]	94.00 [91.00, 96.00]
CRP, mg/L, (median [IQR])	792 (40.8)	157 (44.9)	635 (40.0)
WBC, 10 ⁹ /L, (median [IQR])	79.00 [40.38, 135.00]	82.60 [40.72, 134.62]	78.00 [40.25, 135.00]

PCR positive (%)	1844 (95.7)	314 (89.2)	1530 (97.1)
Time between onset of symptoms and hospital admission, days, (median [IQR])	7.00 [5.00, 12.00]	8.00 [5.00, 13.00]	7.00 [5.00, 12.00]
ICU-admission (%)	348 (17.9)	70 (19.8)	278 (17.4)
In patients admitted to the ICU; days between admission and start of mechanical ventilation *	1.00 [0.00, 3.00]	1.00 [0.00, 3.00]	1.00 [0.00, 3.00]

319

320 (H)CQ denotes (hydroxy)chloroquine; CRP C-reactive protein; WBC white blood cell count; PCR-

321 positive a positive test for COVID-19 based on polymerase chain reaction; ICU intensive care unit. *

322 Data of one centre were missing.

323

324

325 Table 2: Results of Cox-regression models for treatment strategy

326

Cox-regression for treatment strategy

	HR	95%CI		p-value
(H)CQ treatment strategy	1.09	0.81	1.47	0.568
Women	1.04	0.84	1.29	0.715
Age	1.07	1.06	1.08	<0.001
Chronic cardiac disease	1.23	0.98	1.53	0.068
Asthma or chronic pulmonary disease	1.14	0.91	1.42	0.250
Chronic kidney disease (%)	0.99	0.74	1.31	0.919
Malignant neoplasm or chronic hematologic disorder (%)	1.34	1.00	1.79	0.051
Diabetes	1.34	1.07	1.68	0.010
Hypertension	1.06	0.85	1.33	0.577
Obesity	1.23	0.97	1.57	0.087
Peripheral oxygen saturation	0.95	0.94	0.97	<0.001
Respiratory rate	1.04	1.03	1.06	<0.001
Use of steroids for ARDS	1.78	1.26	2.52	0.001
Time in pandemic	0.98	0.97	0.99	<0.001

327 (H)CQ denotes (hydroxy)chloroquine, ARDS acute respiratory distress syndrome. HR indicate

328 multivariable hazard ratios, 95%CI the lowest and highest values of confidence interval.

329

330 Table 3: Results of Cox-regression models for actual treatment

331

Cox-regression for actual treatment

	HR	95%CI		p-value
(H)CQ treatment	1.58	1.24	2.02	<0.001
Women	1.06	0.86	1.31	0.587
Age	1.07	1.06	1.08	0.000
Chronic cardiac disease	1.26	1.01	1.57	0.041
Asthma or chronic pulmonary disease	1.10	0.89	1.37	0.377
Chronic kidney disease (%)	1.00	0.75	1.32	0.977
Malignancy or chronic hematologic disorder (%)	1.36	1.02	1.82	0.037
Diabetes	1.33	1.06	1.66	0.014
Hypertension	1.06	0.85	1.32	0.610
Obesity	1.25	0.98	1.59	0.074
Peripheral oxygen saturation	0.95	0.94	0.97	0.000
Respiratory rate	1.04	1.02	1.06	0.000
Use of steroids for ARDS	1.62	1.14	2.28	0.007
Time in pandemic	0.99	0.98	0.99	0.001

332 (H)CQ denotes (hydroxy)chloroquine, ARDS acute respiratory distress syndrome. HR indicate

333 multivariable hazard ratios, 95%CI the lowest and highest values of confidence interval.

334

335

336 Table 4: Complete cases analysis using inverse probability weighting

Complete case analysis using inverse probability weighting

For treatment strategy

	HR	95%CI		p-value
(H)CQ treatment strategy	1.17	0.99	1.40	0.072

337

338 (H)CQ denotes (hydroxy)chloroquine.

339

340 **Complete case analysis using inverse probability weighting**

341 For actually received treatment

	HR	95%CI		p-value
(H)CQ received treatment	1.41	1.19	1.66	<0.001

342 (H)CQ denotes (hydroxy)chloroquine.

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